

Stat 145, Tue 19-Oct-2021 -- Tue 19-Oct-2021
Biostatistics
Spring 2021

Binary categorical vars?
Is Democrat?
 \bar{X} same as \hat{p} .

Tuesday, October 19th 2021

Due:: WW NormalDists due at 11 pm

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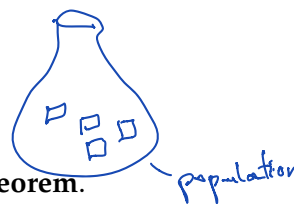
Wk 8, Tu

Topic:: Inference for proportions

Read:: Lock5 6.1-6.3

$$\bar{X} = \frac{1}{n} (X_1 + X_2 + \dots + X_n)$$

Central Limit Theorem 5.2



In summary, here is the take-away from the **Central Limit Theorem**.

Suppose you have a random sample of size n that is either

- i.i.d., or
- an SRS, with the sample size n being no more than 10% of the size of the population.

In the case that

1. the variable under consideration is quantitative, having population mean μ and standard deviation σ , then the sampling distribution for the sample mean \bar{X} is approximately $\text{Norm}(\mu, \sigma / \sqrt{n})$ for n large enough.
2. the variable under consideration is binary categorical, having population proportion p , then the sampling distribution for the sample proportion \hat{p} is approximately $\text{Norm}(p, \sqrt{p(1-p)/n})$ for n large enough.

Since

- null distributions
- randomization distributions
- bootstrap distributions

are all specialized versions of sampling distributions, then so long as the sample statistic in question is the sample's *mean* \bar{X} or the sample *proportion* \hat{p} , we can expect the CLT to apply to these as well.

Explorations using apps at

<https://shiny.calvin.edu:3838/scofield/samplingDists/>

<https://shiny.calvin.edu:3838/scofield/cltProportions/>

or, for means, use script `samplingDistOfSampleMeanExperiments.R`

```
require(mosaic)
require(gridExtra)

# Create a population
mypop <- 50 - rexp(10000, rate=.15)           # left-skewed
#mypop <- rgamma(10000, shape=1.6, rate=.1)  # right-skewed
#mypop <- rnorm(10000, mean=25, sd=6)       # normal
print(favstats(~mypop))

# Simulate the sampling distribution for the sample mean
sampleSize = 20
manyMeans <- do(5000) * mean(~sample(mypop, sampleSize, replace=TRUE))
print(favstats(~mean, data=manyMeans))

p1 <- gf_density(~mypop) %>% gf_refine(scale_x_continuous(limits=c(0,55)))
p2 <- gf_density(~mean, data=manyMeans) %>%
  gf_refine(scale_x_continuous(limits=c(0,55)))
grid.arrange(p1, p2, nrow=2)
```

Code for
experiments

Chapter 6

6.1-6.3: univariate, binary categorical data (single-proportions)

6.4-6.6: univariate, quantitative data (single means)

6.7-6.9: bivariate binary responses for 2 groups (2-proportions)
categorical

6.10-6.12: bivariate data, quantitative response, 2 groups (2 means)

Single proportions: tasks

1. CI for p

2. Hyp. test on null/alt. hypotheses about p

Recall CI in the past

$$\hat{p} \pm z^* (SE_{\hat{p}})$$

↑
sample est.
of p

Note that

for 95% conf., $z^* = 1.96$

90% conf., $z^* = 1.645$

99% conf., $z^* = 2.576$

Ex.] Suppose we ask 450 students if they are left-handed,

47 say yes.

$$\text{So } \hat{p} = \frac{47}{450} = 0.104$$

Q: Can I assume \hat{p} has a nearly normal dist?

$$n\hat{p} = 450 \left(\frac{47}{450} \right) = 47 \geq 10$$

$$n(1-\hat{p}) = 450 \left(\frac{450-47}{450} \right) = 403 \geq 10$$

A: Yes. So $\hat{p} \sim \text{Norm}\left(p, \sqrt{\frac{p(1-p)}{n}}\right)$

$$\begin{aligned} \text{Use } \hat{p}: \text{SE} &= \sqrt{\frac{\hat{p}(1-\hat{p})}{n}} \\ &= \sqrt{\frac{(-.104)(0.896)}{450}} \\ &= 0.0144 \end{aligned}$$

So, a 95% CI for p

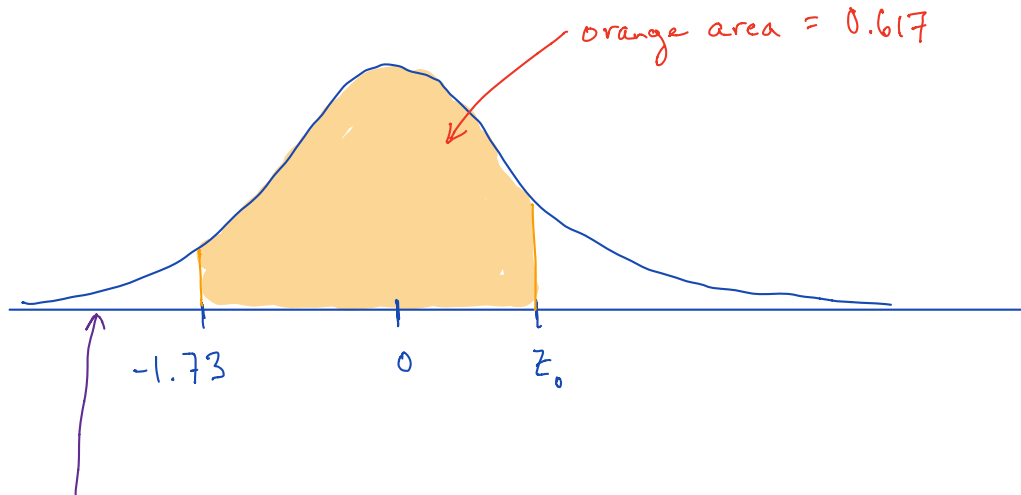
$$0.104 \pm (1.96)(0.0144)$$

↑ ↑ ↑
 \hat{p} z^* $\text{SE}_{\hat{p}}$

A problem like 8(f) on the current WebWork set might be:

"Determine z_0 so that $\Pr(-1.73 < Z < z_0) = 0.617$."

I have depicted a standard normal distribution below with -1.73 , z_0 , and 0.617 displayed.



Use $\text{pnorm}(-1.73)$, which behaves the same as $\text{pnorm}(-1.73, \text{mean}=0, \text{sd}=1)$, to find the area A to the left of 1.73 .

Then use $\text{qnorm}(A + 0.617)$ to find z_0 .

Chapter 6 overview

- Scenarios are all ones we have discussed
 - univariate (one population)
 - proportion arising from binary categorical variable
 - mean arising from quantitative variable
 - 2 populations
 - difference of proportions
 - difference of means investigated using
 - two independent samples
 - matched pairs
- Deferred to later chapter: 2 quant vars
- Can see Chapter 6 as something of a history lesson
- Relies entirely on facts from Central Limit Theorem

Sections 1-3: single proportion

Confidence interval construction

- review how done using bootstrapping (Ch. 3)
- refining the z^* -value
 - in past, stats students used tables of Z-scores
 - see <https://www.math.arizona.edu/~jwatkins/normal-table.pdf>
 - compare with `pnorm()`, `qnorm()` calculations
- formula for SE

Practice:

- obtaining critical z^* values for
 - 96% confidence
 - 90% confidence
 - 99% confidence
- doing inference (CI and hypothesis testing) with datasets
 1. in 119 games of rock-paper-scissors, player did rock 66 times
 2. in 70 out of 120 soccer games, the home team won
 3. suppose that 42% of people have O+ blood. sample shows 65 out of 192