Stat 145, Wed 20-Oct-2021 -- Wed 20-Oct-2021
Biostatistics
Spring 2021

Wednesday, October 20th 2021 Due:: Moodle Quiz Ch. 4 at 11 pm Wednesday, October 20th 2021 Wk 8, We Topic:: Inference for one proportion wrapup Read:: Lock5 6.2-6.3 HW((= PSIC Theorem)

## Some finer points concerning inference on one proportion

**Confidence intervals** for *p*:

• Centered interval approach

This approach makes sense if it's reasonable to assume  $\hat{p}$  has a normal distribution, so check the rules of thumb: ),,,,,,,,

$$n\widehat{p} \ge 10$$
 and  $n(1-\widehat{p}) \ge 10.$   $n(1-p)$  loth at [art [0]

(In CI construction, you have  $\hat{p}$ , but not *p*.)

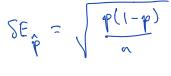
• Formula 
$$SE_{\widehat{p}} = \frac{\sqrt{p(1-p)}}{\sqrt{n}}$$
 is modified to  $SE_{\widehat{p}} = \frac{\sqrt{p(1-\widehat{p})}}{\sqrt{n}}$  when building CI for  $p$ .  
Note that  $SE_{\widehat{p}}$  decreases as  $n$  grows. Specifically.

Note that  $SE_{\hat{p}}$  decreases as *n* grows. Specifically,

- can cut SE<sub> $\hat{v}$ </sub> in half by quadrupling (4×) sample size
- can cut SE<sub> $\hat{v}$ </sub> to one third its size at *n* by increasing sample size by a factor of 9
- other similar statements?

• Can "tailor" a margin of error ME

$$ME = (2^{k})(SE_{p}) = 2^{k} \sqrt{\frac{p(1-p)}{n}}$$
Square loth sides
$$(ME)^{2} = (2^{k})^{2} \frac{p(1-p)}{n}$$
Solve for n:
$$N \approx \left(\frac{2^{k}}{ME}\right)^{2} \frac{p(1-p)}{n}$$
problem? Den't know p
$$Worst case at p = 1/2$$
So if any goal is a ME no bigger than M, should be able to achieve
that by taking somple size
$$N \approx \left(\frac{2^{k}}{M}\right)^{2} \frac{(\frac{2^{k}}{N})^{2}}{2} \frac{(\frac{2^{k}}{N})^{2}}{2} \frac{(\frac{2^{k}}{N})^{2}}{2} \frac{(\frac{2^{k}}{N})^{2}}{2}$$





**Hypothesis tests** involving null hypothesis  $\mathbf{H}_0: p = p_0$ 

- several **test statistics** one might use from the sample
  - 1. the sample proportion  $\hat{p}$ 
    - natural to use, as an unbiased estimator of *p*
    - approximating normal distribution: Norm $(p_0, \sqrt{p_0(1-p_0)/n})$
  - 2. the count of successes X

null value at center

SE

- natural to use, as an unbiased estimator of  $\boldsymbol{p}$
- approximating normal distribution: Norm $(np_0, \sqrt{np_0(1-p_0)})$
- 3. the Z-score of either of the previous, calculated either as

$$z = \frac{\widehat{p} - p_0}{\sqrt{p_0(1 - p_0)/n}}$$
, or  $z = \frac{X - np_0}{\sqrt{np_0(1 - p_0)}}$ .

- called the **standardized test statistic**
- approximating normal distribution: Norm(0, 1)
- one advantage: immediate comparison with critical  $z^*$  value

Practice:

1. Obtain the critical  $z^*$  value for 96% confidence.

gnorm (0.98), or 2\* = 2.054

2. One might expect the 3 options in the "rock, paper, scissors" game are equally likely. If a certain player shows "rock" in 51 out of 119 tries, test that rock appears one-third of the time vs. a 2-sided alternative. Use significance level alpha = 0.04.
P = proportion of "rock"

H<sub>0</sub>: 
$$p = \frac{1}{3}$$
 rs. H<sub>a</sub>:  $p \neq \frac{1}{3}$   
 $p = \frac{51}{119} = 0.4286$  Estanted SE:  $\sqrt{\frac{p(1-p)}{n}} = \sqrt{\frac{(\frac{1}{3}\sqrt{2}/3)}{119}} = 0.04381$   
Can (define 1) consult Norm ( $\frac{1}{3}$ , 0.04321) or (define 2) standardize the test statistic  
 $Z = \frac{0.4286 - \frac{1}{3}}{0.04321} = 2.205$   
and consult Normal(0,1)  
(3.) If you wish a 95% CI for p to have a margin of error around 0.03,  $n \geq \left(\frac{2^{*}}{2^{*}}\right)^{2}$   
M=0.63,  $2^{*} = 1.96$   
 $n \geq \left(\frac{1.96}{2(6.03)}\right)^{2} = 1067.1$ 

4. What might be the minimum sample size n in the last question if you have the additional information that p cannot be larger than 0.25? 2

$$n \ge \left(\frac{Z^{*}}{M}\right)^{2} p(1-p) = \left(\frac{1.96}{0.03}\right)^{2} (0.25)(1-0.25) =$$

Take a at least 1068,